

## Truncated Moments of the Reciprocal of a Noncentral F Variable

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### Abstract

The bias and risk of the positive part Stein type estimators involve the evaluation of truncated moments of the reciprocal of noncentral F-variables (Saleh 2004). Bock et al (1984) obtained simple form for the inverse moments of non-central  $\chi^2$  and  $F$  random variables, which depend on the Dawson's integral (Abramowitz and Stegun 1964). However, they have not indicated about the truncated moments for the reciprocal of noncentral  $F$  variables. In this note, we find the truncated moments of the reciprocal of a noncentral F variable in terms of incomplete beta function.

**Keywords and Phrases:** Truncated moment; Non-central F; Incomplete Beta.

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### Inverse moments of a truncated noncentral F random variable

Suppose

$$Z_1 \sim \chi_{n_1, \Delta}^2 \quad \text{and} \quad Z_2 \sim \chi_{n_2}^2,$$

where  $Z_1$  and  $Z_2$  are mutually independent. Then

$$X = \frac{Z_1/n_1}{Z_2/n_2}$$

is distributed as noncentral F-distribution with  $n_1, n_2$  degrees of freedom and non-centrality parameter  $\Delta$ . The *pdf* of  $X$  can be written as a function of central F and a

power series

$$f(x; \Delta) = f(x)e^{-\Delta/2} \sum_{j=0}^{\infty} \left( \frac{\frac{1}{2}\Delta n_1 x}{n_2 + n_1 x} \right)^j h(n_j), \quad (1)$$

where  $f(x)$  is the central F distribution with  $(n_1, n_2)$  degrees of freedom and

$$h(n_j) = \frac{(n_1 + n_2)(n_1 + n_2 + 2)...(n_1 + n_2 + 2\overline{j-1})}{j! n_1 (n_1 + 2)...(n_1 + 2\overline{j-1})}.$$

The *cdf* of noncentral  $F$ , which can also be expressed as incomplete noncentral beta function ratio as

$$\begin{aligned} P(x_{n_1, n_2}(\Delta) \leq f) &= I_{\left(\frac{n_1 f}{n_2 + n_1 f}\right)} \left( \frac{n_1}{2}, \frac{n_2}{2}; \Delta \right) \\ &= \sum_{j=0}^{\infty} \left( \frac{(\frac{1}{2}\Delta)^j e^{-\Delta/2}}{j!} \right) I_{\left(\frac{n_1 f}{n_2 + n_1 f}\right)} \left( \frac{n_1}{2} + j, \frac{n_2}{2} \right), \end{aligned} \quad (2)$$

where  $I_p(\alpha, \beta)$  is the incomplete beta function ratio given by

$$I_p(\alpha, \beta) = \frac{1}{B(\alpha, \beta)} \int_0^p t^{\alpha-1} (1-t)^{\beta-1} dt.$$

Our interest is to find the truncated  $r^{th}$  moment of the reciprocal of a noncentral F variable and this will be done by induction method.

*Case 1:* For,  $r = 1$ , to find  $E(x^{-1} I_{x \leq d_1})$ , note that

$$\begin{aligned} E(x^{-1} I_{x \leq d_1}) &= \int_0^{d_1} x^{-1} \frac{(\frac{n_1}{n_2})^{n_1/2}}{B(\frac{n_1}{2}, \frac{n_2}{2})} \frac{x^{\frac{n_1}{2}-1}}{(1 + \frac{n_1}{n_2}x)^{\frac{n_1+n_2}{2}}} e^{-\Delta/2} \sum_{j=0}^{\infty} \left( \frac{\frac{1}{2}\Delta n_1 x}{n_2 + n_1 x} \right)^j h(n_j) dx \\ &= \int_0^{d_1} \frac{(\frac{n_1}{n_2})^{n_1/2}}{B(\frac{n_1}{2}, \frac{n_2}{2})} \frac{x^{\frac{n_1-2}{2}-1}}{(1 + \frac{n_1}{n_2}x)^{\frac{n_1+n_2}{2}}} e^{-\Delta/2} \sum_{j=0}^{\infty} \left( \frac{\frac{1}{2}\Delta n_1 x}{n_2 + n_1 x} \right)^j h(n_j) dx. \end{aligned}$$

Let us make the following transformation

$$\frac{n_1}{n_2}x = \frac{n_1 - 2}{n_2 + 2}y \Rightarrow x = \frac{n_2(n_1 - 2)}{n_1(n_2 + 2)}y \Rightarrow dx = \frac{n_2(n_1 - 2)}{n_1(n_2 + 2)}dy$$

If  $x = 0$ , then  $y = 0$ , if  $x = d_1$ ,  $y = d_1^* = \frac{n_1(n_2+2)}{n_2(n_1-2)}$ . We note after simple calculation that,

$$\frac{\frac{1}{2}\Delta n_1 x}{n_2 + n_1 x} = \frac{\frac{1}{2}\Delta(n_1 - 2)y}{(n_2 + 2) + (n_1 - 2)y}.$$

We have,

$$\begin{aligned}
E(x^{-1} I_{x \leq d_1}) &= \int_0^{d_1^*} \frac{\left(\frac{n_1}{n_2}\right)^{n_1/2}}{B\left(\frac{n_1}{2}, \frac{n_2}{2}\right)} \frac{\left(\frac{n_2(n_1-2)}{n_1(n_2+2)}y\right)^{\frac{n_1-2}{2}-1}}{(1 + \frac{n_1-2}{n_2+2}y)^{\frac{n_1+n_2}{2}}} \frac{n_2(n_1-2)}{n_1(n_2+2)} \\
&\times e^{-\Delta/2} \sum_{j=0}^{\infty} \left( \frac{\frac{1}{2}\Delta(n_1-2)y}{(n_2+2) + (n_1-2)y} \right)^j h(n_j) dy \\
&= \int_0^{d_1^*} \frac{\left(\frac{n_1}{n_2}\right) \left(\frac{(n_1-2)}{(n_2+2)}\right)^{\frac{n_1-2}{2}}}{B\left(\frac{n_1}{2}, \frac{n_2}{2}\right)} \frac{(y)^{\frac{n_1-2}{2}-1}}{(1 + \frac{n_1-2}{n_2+2}y)^{\frac{n_1+n_2}{2}}} e^{-\Delta/2} \\
&\times \sum_{j=0}^{\infty} \left( \frac{\frac{1}{2}\Delta(n_1-2)y}{(n_2+2) + (n_1-2)y} \right)^j h(n_j) dy. \tag{3}
\end{aligned}$$

Since

$$\frac{\left(\frac{n_1}{n_2}\right)}{B\left(\frac{n_1}{2}, \frac{n_2}{2}\right)} = \frac{\left(\frac{n_1}{n_2-2}\right)}{B\left(\frac{n_1-2}{2}, \frac{n_2+2}{2}\right)},$$

the equation (3) can be written as

$$\begin{aligned}
E(x^{-1} I_{x \leq d_1}) &= \int_0^{d_1^*} \frac{\left(\frac{n_1}{n_1-2}\right) \left(\frac{(n_1-2)}{(n_2+2)}\right)^{\frac{n_1-2}{2}}}{B\left(\frac{n_1-2}{2}, \frac{n_2+2}{2}\right)} \frac{(y)^{\frac{n_1-2}{2}-1}}{(1 + \frac{n_1-2}{n_2+2}y)^{\frac{n_1+n_2}{2}}} \\
&\times e^{-\Delta/2} \sum_{j=0}^{\infty} \left( \frac{\frac{1}{2}\Delta(n_1-2)y}{(n_2+2) + (n_1-2)y} \right)^j h(n_j) dy \\
&= n_1 \int_0^{d_1^*} \frac{\left(\frac{(n_1-2)}{(n_2+2)}\right)^{\frac{n_1-2}{2}}}{B\left(\frac{n_1-2}{2}, \frac{n_2+2}{2}\right)} \frac{(y)^{\frac{n_1-2}{2}-1}}{(1 + \frac{n_1-2}{n_2+2}y)^{\frac{n_1+n_2}{2}}} e^{-\Delta/2} \\
&\times \sum_{j=0}^{\infty} \left( \frac{\frac{1}{2}\Delta(n_1-2)y}{(n_2+2) + (n_1-2)y} \right)^j h_1(n_j) dy \\
&= n_1 \int_0^{d_1^*} f(y, n_1-2, n_2+2) e^{-\Delta/2} \\
&\times \sum_{j=0}^{\infty} \left( \frac{\frac{1}{2}\Delta(n_1-2)y}{(n_2+2) + (n_1-2)y} \right)^j h_1(n_j) dy \\
&= n_1 P(Y_{n_1-2, n_2+2}(\Delta) \leq d_1^*), \tag{4}
\end{aligned}$$

where  $f(y, n_1 - 2, n_2 + 2)$  is the *pdf* of central F distribution with  $(n_1 - 2, n_2 + 2)$  degrees of freedom,  $Y_{n_1-2, n_2+2}(\Delta)$  has a noncentral F distribution with  $(n_1 - 2, n_2 + 2)$  degrees of freedom and noncentrality parameter  $\Delta$ , and

$$h_1(n_j) = \frac{(n_1 + n_2)(n_1 + n_2 + 2) \dots (n_1 + n_2 + 2\bar{j} - 1)}{j!(n_1 - 2)n_1(n_1 + 2) \dots (n_1 - 2 + 2\bar{j} - 1)}.$$

Now comparing equations (2) and (4), we have

$$E(x^{-1}I_{x \leq d_1}) = n_1 \sum_{j=0}^{\infty} \frac{(\frac{\Delta}{2})^j e^{-\Delta/2}}{j!} I_{\left(\frac{(n_1-2)d_1^*}{n_2+2+(n_1-1)d_1^*}\right)} \left(\frac{n_1-2}{2} + j, \frac{n_2+2}{2}\right).$$

Since

$$\frac{(n_1-2)d_1^*}{n_2+2+(n_1-1)d_1^*} = \frac{n_1d_1}{n_2+n_1d_1},$$

we get

$$\begin{aligned} E(x^{-1}I_{x \leq d_1}) &= n_1 \sum_{j=0}^{\infty} \frac{(\frac{\Delta}{2})^j e^{-\Delta/2}}{j!} I_{\left(\frac{n_1d_1}{n_2+n_1d_1}\right)} \left(\frac{n_1-2}{2} + j, \frac{n_2+2}{2}\right) \\ &= n_1 \times I_{\left(\frac{n_1d_1}{n_2+n_1d_1}\right)} \left(\frac{n_1-2}{2}, \frac{n_2+2}{2}; \Delta\right). \end{aligned} \quad (5)$$

*Case 2:* For,  $r = 2$ , to find  $E(x^{-2}I_{x \leq d_2})$ , note that

$$\begin{aligned} E(x^{-2}I_{x \leq d_2}) &= \int_0^{d_2} x^{-2} \frac{(\frac{n_1}{n_2})^{n_1/2}}{B(\frac{n_1}{2}, \frac{n_2}{2})} \frac{x^{\frac{n_1}{2}-1}}{(1 + \frac{n_1}{n_2}x)^{\frac{n_1+n_2}{2}}} e^{-\Delta/2} \sum_{j=0}^{\infty} \left(\frac{\frac{1}{2}\Delta n_1 x}{n_2 + n_1 x}\right)^j h(n_j) dx \\ &= \int_0^{d_2} \frac{(\frac{n_1}{n_2})^{n_1/2}}{B(\frac{n_1}{2}, \frac{n_2}{2})} \frac{x^{\frac{n_1-4}{2}-1}}{(1 + \frac{n_1}{n_2}x)^{\frac{n_1+n_2}{2}}} e^{-\Delta/2} \sum_{j=0}^{\infty} \left(\frac{\frac{1}{2}\Delta n_1 x}{n_2 + n_1 x}\right)^j h(n_j) dx \end{aligned}$$

Let us make the following transformation

$$\frac{n_1}{n_2}x = \frac{n_1-4}{n_2+4}y \Rightarrow x = \frac{n_2(n_1-4)}{n_1(n_2+4)}y$$

If  $x = 0$ , then  $y = 0$ , if  $x = d_2$ ,  $y = d_2^* = \frac{n_1(n_2+4)}{n_2(n_1-4)}$ . Also,

$$\frac{\frac{1}{2}\Delta n_1 x}{n_2 + n_1 x} = \frac{\frac{1}{2}\Delta(n_1-4)y}{(n_2+4) + (n_1-4)y}.$$

Thus

$$E(x^{-2}I_{x \leq d_2}) = \int_0^{d_2^*} \frac{(\frac{n_1}{n_2})^{n_1/2}}{B(\frac{n_1}{2}, \frac{n_2}{2})} \frac{\left(\frac{n_2(n_1-4)}{n_1(n_2+4)}y\right)^{\frac{n_1-4}{2}-1}}{(1 + \frac{n_1-4}{n_2+4}y)^{\frac{n_1+n_2}{2}}} \frac{n_2(n_1-4)}{n_1(n_2+4)}$$

$$\times e^{-\Delta/2} \sum_{j=0}^{\infty} \left( \frac{\frac{1}{2}\Delta(n_1-4)y}{(n_2+4)+(n_1-4)y} \right)^j h(n_j) dy. \quad (6)$$

Since

$$\frac{\left(\frac{n_1}{n_2}\right)^2}{B\left(\frac{n_1}{2}, \frac{n_2}{2}\right)} = \frac{\frac{n_1^2(n_2+2)}{n_2(n_1-2)(n_1-4)}}{B\left(\frac{n_1-4}{2}, \frac{n_2+4}{2}\right)},$$

equation (6) can be written as

$$\begin{aligned} E(x^{-2} I_{x \leq d_2}) &= \frac{n_1^2(n_2+2)}{n_2} \int_0^{d_2^*} \frac{\left(\frac{(n_1-4)}{(n_2+4)}\right)^{\frac{n_1-4}{2}}}{B\left(\frac{n_1-4}{2}, \frac{n_2+4}{2}\right)} \frac{(y)^{\frac{n_1-4}{2}-1}}{(1+\frac{n_1-4}{n_2+4}y)^{\frac{n_1+n_2}{2}}} \\ &\times e^{-\Delta/2} \sum_{j=0}^{\infty} \left( \frac{\frac{1}{2}\Delta(n_1-4)y}{(n_2+4)+(n_1-4)y} \right)^j h(n_j) dy \\ &= \frac{n_1^2(n_2+2)}{n_2} \int_0^{d_2^*} f(y, n_1-4, n_2+4) e^{-\Delta/2} \\ &\times \sum_{j=0}^{\infty} \left( \frac{\frac{1}{2}\Delta(n_1-4)y}{(n_2+4)+(n_1-4)y} \right)^j h_2(n_j) dy \\ &= \frac{n_1^2(n_2+2)}{n_2} P(Y_{n_1-4, n_2+4}(\Delta) \leq d_2^*), \end{aligned} \quad (7)$$

where  $f(y, n_1-4, n_2+4)$  is the pdf of central F distribution with  $(n_1-4, n_2+4)$  degrees of freedom,  $Y_{n_1-4, n_2+4}(\Delta)$  has a noncentral F distribution with  $(n_1-4, n_2+4)$  degrees of freedom, and noncentrality parameter  $\Delta$  and

$$h_2(n_j) = \frac{(n_1+n_2)(n_1+n_2+2)\dots(n_1+n_2+2(j-1))}{j!(n_1-4)(n_1-2)n_1(n_1+2)\dots(n_1-4+2(j-1))}.$$

Now comparing equations (2) and (7), we have

$$E(x^{-2} I_{x \leq d_2}) = \frac{n_1^2(n_2+2)}{n_2} \sum_{j=0}^{\infty} \frac{\left(\frac{\Delta}{2}\right)^j e^{-\Delta/2}}{j!} I_{\left(\frac{(n_1-4)d_2^*}{n_2+4+(n_1-4)d_2^*}\right)} \left( \frac{n_1-4}{2} + j, \frac{n_2+4}{2} \right).$$

Since

$$\frac{(n_1-4)d_2^*}{n_2+4+(n_1-4)d_2^*} = \frac{n_1 d_2}{n_2 + n_1 d_2},$$

we have

$$E(x^{-2} I_{x \leq d_2}) = \frac{n_1^2(n_2+2)}{n_2} \sum_{j=0}^{\infty} \frac{\left(\frac{\Delta}{2}\right)^j e^{-\Delta/2}}{j!} I_{\left(\frac{n_1 d_2}{n_2 + n_1 d_2}\right)} \left( \frac{n_1-4}{2} + j, \frac{n_2+4}{2} \right)$$

$$= \frac{n_1^2(n_2+2)}{n_2} \times I_{\left(\frac{n_1d_2}{n_2+n_1d_2}\right)} \left( \frac{n_1-4}{2}, \frac{n_2+4}{2}; \Delta \right). \quad (8)$$

*Case 3:* For,  $r = 3$ , to find  $E(x^{-3}I_{x \leq d_3})$ , following all steps in cases (1) and (2), we have

$$\begin{aligned} E(x^{-3}I_{x \leq d_3}) &= \frac{n_1^3(n_2+2)(n_2+4)}{n_2^2} \sum_{j=0}^{\infty} \frac{(\frac{\Delta}{2})^j e^{-\Delta/2}}{j!} I_{\left(\frac{n_1d_3}{n_2+n_1d_3}\right)} \left( \frac{n_1-6}{2} + j, \frac{n_2+6}{2} \right) \\ &= \frac{n_1^3(n_2+2)(n_2+4)}{n_2^2} \times I_{\left(\frac{n_1d_3}{n_2+n_1d_3}\right)} \left( \frac{n_1-6}{2}, \frac{n_2+6}{2}; \Delta \right). \end{aligned} \quad (9)$$

Finally by induction, the  $r^{th}$  truncated moment of the reciprocal of a noncentral F variable in terms of incomplete beta function ratio is given by

$$\begin{aligned} E(x^{-r}I_{x \leq d_r}) &= \frac{n_1^r(n_2+2)(n_2+4) \cdots (n_2+2.r-1)}{n_2^{r-1}} \\ &\quad \times \sum_{j=0}^{\infty} \frac{(\frac{\Delta}{2})^j e^{-\Delta/2}}{j!} I_{\left(\frac{n_1d_r}{n_2+n_1d_r}\right)} \left( \frac{n_1-2r}{2} + j, \frac{n_2+2r}{2} \right) \\ &= \left( \frac{n_1}{n_2} \right)^r \prod_{i=1}^r [n_2 + 2(i-1)] I_{\left(\frac{n_1d_r}{n_2+n_1d_r}\right)} \left( \frac{n_1-2r}{2}, \frac{n_2+2r}{2}; \Delta \right). \end{aligned} \quad (10)$$

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### References

- [1] Abramowitz, M. and I. Stegun (1964). Handbook of mathematical functions with formulas, graphs, and mathematical tables (U.S. Department of Commerce, Washington, DC).
- [2] Bock, M. E., Judge, G. G. and Yancey, T. A. (1984). A simple form for the inverse moments of non-central  $\chi^2$  and F random variables and certain confluent Hypergeometric functions. *Journal of Econometrics*, 25, 217-234.
- [3] Saleh, A. K. Md. E. (2004). Theory of preliminary test and Stein-type estimation with applications. A research book for Wiley in the process of publication.