ISSN 1683-5603

International Journal of Statistical Sciences Vol. 3 (Special Issue), 2004, pp 9–15 © 2004 Dept. of Statistics, Univ. of Rajshahi, Bangladesh

# **Glimpses of Inequalities in Probability and Statistics**

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[Received March 12, 2004; Accepted June 16, 2004]

## Abstract

In this article we compile all basic and well-known inequalities useful in probability and statistics. Proofs are given only for two less known inequalities.

**Keywords and Phrases:** Probability inequalities, Moment inequalities, Quantiles, Harmonic mean.

AMS Classification: 60-02, 60E15, 60G50, 62-02, 62E17.

# 1 Introduction

Rao (1996) wrote an article entitled, "Seven inequalities in statistical estimation theory", that has appeared in the journal, Student[6]. The present article is almost a complement of Professor Rao's article. Professor Rao's inequalities were on estimation theory, whereas our inequalities are primarily on probability theory. Most of the inequalities are well-known, but some of the inequalities may not be known to all. This compilation of basic inequalities is expected to serve as reference to most researches in probability and statistics.

## 2 Basic Inequalities

In this section we state fifteen basic well-known inequalities.

#### 2.1 Markov's Inequality.

Suppose that X is a nonnegative random variable with finite mean, E(X). Then for any t > 0,

$$P(X \ge t) \le \frac{E(X)}{t}.$$

## 2.2 Chebyshev's Inequality.

Suppose that X is a random variable with finite mean  $\mu$  and variance  $\sigma^2$ . Then for any t > 0,

or

$$P\left(|X - \mu| \ge t\right) \le \frac{\sigma^2}{t^2}$$
$$P\left(|X - \mu| < t\right) \ge 1 - \frac{\sigma^2}{t^2}.$$

### 2.3 One-sided Chebyshev's inequality.

Let X be a random variable with mean  $\mu$  and variance  $\sigma^2$ . Then for any t > 0,

(i) 
$$P(X \ge \mu + t) \le \frac{\sigma^2}{\sigma^2 + t^2}$$
  
(ii)  $P(X \le \mu - t) \le \frac{\sigma^2}{\sigma^2 + t^2}$ .

## 2.4 Chernoff bounds.

Let X be any r.v. with moment generating function  $M_X(t)$ . Then,

(i) 
$$P(X \ge s) \le \frac{M_X(t)}{e^{ts}}$$
 for all  $t > 0$   
(ii)  $P(X \le s) \le \frac{M_X(t)}{e^{ts}}$  for all  $t < 0$ .

#### 2.5 Basic Inequality.

Let g be a nonnegative and nondecreasing function on R on the range of the random variable X. Further assume that ess.sup  $g(X) < \infty$ . Then

$$\frac{E\{g(X)\} - g(a)}{\text{ess. sup } g(X)} \le P(X \ge a) \le \frac{E\{g(X)\}}{g(a)}$$

In addition, if g is an even function, replace  $P(X \ge a)$  by  $P(|X| \ge a)$  in the above. Here ess. sup g(x) is the least among constants below which almost all values of g(x) lie.

#### 2.6 Corollary to the Basic Inequality.

Let  $g(X) = |X|^r$  for r > 0. Then by basic inequality, we have

$$\frac{E(|X|^r) - a^r}{ess.sup\,g(X)} \le P(|X| \ge a) \le \frac{E(|X|^r)}{a^r}.$$

(i) For r = 1 right hand side becomes the Markov's inequality.

(ii) For r = 2 right hand side becomes the Chebyshev's inequality.

#### 2.7Cauchy-Schwarz Inequality.

Let X and Y be any two random variables. Then

$$[E(XY)]^2 \le E(X^2)E(Y^2),$$

or equivalently,

$$|E(XY)| \le [E(X^2)]^{\frac{1}{2}} [E(Y^2)]^{\frac{1}{2}}.$$

#### $\mathbf{2.8}$ $C_r$ -inequality.

Let X and Y be any two random variables. Then for  $r \ge 0$ ,

$$E\{|X+Y|^r\} \le C_r[E\{|X|^r\} + E\{|X|^r\}],\$$

where

$$C_r = \begin{cases} 1 & \text{for } r \le 1\\ 2^{r-1} & \text{for } r > 1. \end{cases}$$

#### Holder's inequality. $\mathbf{2.9}$

Let X and Y be any two random variables. Then for r > 1, s > 1 with  $\frac{1}{r} + \frac{1}{s} = 1$ 

$$E\{|XY|\} \le [E(|X|^r)]^{\frac{1}{r}} [E(|Y|^s)]^{\frac{1}{s}}.$$

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#### 2.10Lyapounov's inequality.

For a random variable X

$$\left(E(|X|^{\alpha})\right)^{1/\alpha} \leq \left(E(|X|^{\beta})\right)^{1/\beta}, \ 0 < \alpha \leq \beta.$$

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#### 2.11 Minkowski Inequality.

Let X and Y be any two random variables. Then for  $r \ge 1$ ,

$$[E(|X+Y|^r)]^{1/r} \le [E(|X|^r)]^{1/r} + [E(|Y|^r)]^{1/r}.$$

#### 2.12 Jensen's inequality.

Let g be a convex function and E(X) be finite. Then

$$E(g(X)) \ge g(E(X)).$$

### **2.13** An Inequality on the Expectation of $E|X|^{\nu}$ .

Let X be a random variable such that  $E|X|^{\beta} < \infty$  for some  $\beta > 0$ . Then for  $0 \le \alpha \le \beta$ 

$$E|X|^{\alpha} \le 1 + E|X|^{\beta}$$

and

 $E|X|^{\alpha} < \infty.$ 

#### 2.14 Bernstein Inequality (Binomial case).

Let  $S_n$  follow binomial distribution with parameters n and p. Then for  $\epsilon > 0$ 

$$P(|S_n - np| \ge n\epsilon) \le 2e^{-n\epsilon^2/2}.$$

That is, the probability that  $S_n$  exceeds its expected value np more than a multiple  $n\epsilon$  of n, converges to zero exponentially fast as  $n \to \infty$ . Here  $S_n$  may also be viewed as the sums of n independent Bernoulli trials.

#### 2.15 Hoeffding's Inequality.

We propose to give a more general version of this important inequality. Let  $X_1, X_2, \dots, X_n$  be independently distributed random variables with finite means  $E(X_i) = \mu_i$  and  $a_i \leq X_i \leq b_i$  for all *i*. Also, let  $S_n = \sum_{i=1}^n X_i$ . Then for  $\epsilon > 0$ 

$$P(|S_n - E(S_n)| \ge n\epsilon) \le 2e^{-2n^2\epsilon^2/\sum_{i=1}^n (b_i - a_i)^2}.$$

# 3 Further Inequalities

In this section we state two additional less known inequalities on quantiles and the expectation of harmonic mean. In both cases we supply the proofs.

#### 3.1 An Inequality on Quantiles.

Let  $\zeta_p$  be the *p*th quantile of the random variable X that has mean  $\mu$  and variance  $\sigma^2$ . Then

$$\mu - \sigma \sqrt{\frac{1-p}{p}} \le \zeta_p \le \mu + \sigma \sqrt{\frac{p}{1-p}}.$$

For  $0 , the quantity <math>\zeta_p$  is said to be the *p*th quantile of the random variable X if  $P(X \leq \zeta_p) \geq p$  and  $P(X \geq \zeta_p) \geq 1 - p$ .

**Proof.** By definition,  $\zeta_p$  satisfies the inequality

$$p \le P(X \le \zeta_p) = P\left(\frac{X-\mu}{\sigma} \le \frac{\zeta_p-\mu}{\sigma}\right) = P(Z \le t),$$

where  $Z = \frac{X-\mu}{\sigma}$  which has mean 0 and variance 1, and  $t = \frac{\zeta_p - \mu}{\sigma}$ . Now for  $\zeta_p < \mu$ , we have t < 0, and consequently using one-sided Chebyshev's inequality on  $P(Z \le t)$  we have

$$p \le P(Z \le t) \le \frac{1}{1+t^2},$$

i.e.

$$p \le \frac{1}{1+t^2} = \frac{1}{1+\left(\frac{\xi_p-\mu}{\sigma}\right)^2}$$

i.e.

$$-\sqrt{\frac{1-p}{p}} \le \frac{\zeta_p - \mu}{\sigma}$$
$$\mu - \sigma \sqrt{\frac{1-p}{p}} \le \zeta_p.$$

Again by definition,  $\zeta_p$  satisfies the inequality

$$1 - p \le P(X \ge \zeta_p) = P\left(\frac{X - \mu}{\sigma} \ge \frac{\zeta_p - \mu}{\sigma}\right) = P(Z \ge t).$$

Now for  $\zeta_p > \mu$ , we have t > 0, and consequently using one-sided Chebyshev's inequality on  $P(Z \ge t)$  we have

$$1 - p \le P(Z \ge t) \le \frac{1}{1 + t^2},$$

i.e.

$$1-p \leq \frac{1}{1+t^2} = \frac{1}{1+\left(\frac{\xi_p-\mu}{\sigma}\right)^2}$$

i.e.

$$\frac{\zeta_p - \mu}{\sigma} \le \sqrt{\frac{p}{1 - p}}$$

or equivalently

$$\zeta_p \le \mu + \sigma \sqrt{\frac{1-p}{p}}$$

Now combining both limits, we have

$$\mu - \sigma \sqrt{\frac{1-p}{p}} \le \zeta_p \le \mu + \sigma \sqrt{\frac{p}{1-p}}.$$

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**Corollary.** The lower and upper bounds on  $\zeta_{\frac{1}{2}}$  (median) are given by

$$\mu - \sigma \le \zeta_{\frac{1}{2}} \le \mu + \sigma.$$

# 3.2 An Inequality on the Expectation of Harmonic Mean (see Rao, 1996).

Let X > 0 and Y > 0 be random variables. Then

$$E\left(\frac{XY}{X+Y}\right) \le \frac{E(X)E(Y)}{E(X+Y)}.$$

**Proof** (in a special case). Let  $\frac{E(X)}{E(Y)} = \frac{r}{s}$  be rational, and take  $X_1 = \frac{X}{r}$ ,  $Y_1 = \frac{Y}{s}$  so that  $E(X_1) = E(Y_1) = \mu > 0$ . Since geometric mean is greater than or equal to harmonic mean, thus we have

$$\frac{r+s}{\frac{r}{Y_1}+\frac{s}{X_1}} \le Y_1^{r/(r+s)} X_1^{s/(r+s)}.$$

Now taking expectations on both sides we get

$$E\left(\frac{r+s}{\frac{r}{Y_1}+\frac{s}{X_1}}\right) \le E\left(Y_1^{r/(r+s)}X_1^{s/(r+s)}\right).$$

Note that  $\frac{XY}{X+Y} = \frac{rs}{r+s} \cdot \frac{r+s}{\frac{r}{Y_1} + \frac{s}{X_1}}$ . Thus we have

$$E\left(\frac{XY}{X+Y}\right) = \frac{rs}{r+s} E\left(\frac{r+s}{\frac{r}{Y_1} + \frac{s}{X_1}}\right) \le \frac{rs}{r+s} E\left(Y_1^{r/(r+s)} X_1^{s/(r+s)}\right).$$

Using Holder's inequality, we obtain

$$E\left(Y_1^{r/(r+s)}X_1^{s/(r+s)}\right) \le (E(Y_1))^{r/(r+s)} (E(X_1))^{s/(r+s)} = \mu.$$

Finally, we establish

$$E\left(\frac{XY}{X+Y}\right) \le \frac{rs}{r+s}\mu = \frac{E(X)E(Y)}{E(X+Y)}.$$

# 4 Concluding remarks

The inequalities given in this article are useful for solving problems and proving results. These inequalities are a useful tool even for solving research problems. The proofs of the inequalities in section 2 are omitted since these are available in standard text books cited below. The collection of all these inequalities in a compact form becomes handy to students and teachers for reference purposes.

#### Acknowledgements.

The authors wish to thank the Editor and a referee for their constructive suggestions on the earlier version of the article which led to the present improved version.

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