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# Adjusting Measurement Errors in Recumbent Length at Birth by Inverse Regression

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#### Abstract

Attempts are made to adjust recumbent length at birth which is assumed to be subject to measurement errors. Inverse regression estimate of recumbent length is found to be free from measurement errors empirically. But sum of squared distance of inverse regression estimates from the observed values is found to be very high. To reduce the sum of squared distance between the observed and estimated recumbent length at birth, arithmetic mean and geometric mean of inverse regression estimate and that of Pervin (2004) estimate are computed that give nearly identical results and capable of removing measurement errors which Pervin's estimate does not.

**Keywords and Phrases:** Inverse Regression, Instrumental Variable, Recumbent Length.

**AMS Classification:**  $62J_{xx}$ .

# 1 Introduction

Correlated variables, say X and Y, are functionally related as

y = f(x). The simplest explicit relationship under ordinary least square setup may be written as

$$Y_i = \beta_0 + \beta_1 X_i + \epsilon_i, \ \epsilon_i \to NID(0, \sigma_\epsilon^2), Cov(X, \epsilon) = 0 \ i = 1, 2, \dots, n$$
(1)

There also exists an inverse function

$$x = f^{-1}(y) \tag{2}$$

When we fail to observe X correctly but with some errors u, so that the errorneous observed value is  $X^0 = X + u$ .

True values of X may be estimated from the estimated explicit functional relationship (1) using (2) which may reduce the amount of measurement error in X. Wright (1928) introduced this concept and Kructchkoff (1967) first applied this concept successfully in the Calibration of pressure Gauge. In earlier concepts, linear functional relationships are considered. Recently, Lavagnini and Magno (2007) provided an overview of univariate calibration by inverse regression where both linear and quadratic forms of (1) are discussed in the set up of ordinary least squares method, that is, the explanatory variable X is free from error and the error term  $\varepsilon$  governing the relationship (1) is uncorrelated with X. The OLS method of estimating the explicit relationship provided by (1) breaks down providing biased inconsistent estimates of underlying parameters if X is subject to measurement error (Fuller, 1987: Cheng et.al, 1998: Wikipedia, 2012). Whitemore (1989) provided an estimator  $X_i^{\star}$  for  $X_i$ based on reliability index. But this estimator is also biased but consistent with smallest mean squared error (Saleh, 2006). In this case, we may use instrumental variable method to estimate the explicit functional relationship (1) for getting unbiased and consistent estimates. Details of univariate and multivariate linear estimation of (1) by instrumental variable method is available in Wikipedia (2012). The problem arises if the explicit form of (1) is quadratic in nature and X is subject to measurement errors. Instrumental variable method is not available for this problem and we have to search for an alternative method.

## 2 Two Stage Method

Let the explicit relationship of Y and X be quadratic as

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 X_i^2 + \epsilon_i; \ \epsilon_i \sim NID(0, \sigma_\epsilon^2), \ Cov(X, \epsilon) \neq 0; \quad X_i^0 = X_i + u_i \quad (3)$$

Here, any choice of instrumental variable z fails to estimate  $\beta_1$  and  $\beta_2$  simultaneously. Let us first consider the linear portion of the model (3) as

$$Y_i = \beta_0 + \beta_1 X_i + \epsilon_i \tag{4}$$

An unbiased and consistent estimator of  $\beta_1$  say  $b_1$  is available from (4) by applying instrumental variable method. Now, we may rewrite (3) as

$$(Y_i - b_1 X_i^0) = \beta_0 + \beta_2 X_i^2 + u_i, \quad u_i \sim NID(0, \sigma_u^2)$$
(5)

Now an unbiased and consistent estimator of  $\beta_2$ , say  $b_2$ , is available from (5) by using the same instrumental variable set as in (4). Finally, we may estimate  $\beta_0$  from (3), if  $b_2$  is significant as,

$$b_0 = \hat{\beta}_0 = \bar{Y} - b_1 \bar{X}^0 - b_2 \bar{X}^{02} \tag{6}$$

The full estimated model (3) is now

$$\hat{Y}_i = b_0 + b_1 X_i + b_2 X_i^2 \tag{7}$$

Now using the inverse regression, we may estimate  $X_i$  as

$$\hat{X}_i = \frac{-b_1 \pm \sqrt{b_1^2 - 4b_2(b_0 - Y_i)}}{2b_2} \tag{8}$$

Of the two values of  $\hat{X}_i$ , we may consider the one closer to  $X_i^0$ . Thus, we may obtain a new set of observations  $(\hat{X}_i, Y_i)$ , i=1,2,...,n. Fresh estimation may be proceeded with  $(\hat{X}_i, Y_i)$ , i=1,2,...,n. The process is repeated until the estimates of parameters converge. From (5),  $b_2$  is asymptotically normal with mean $\beta_2$  and variance  $s^2(z'x^2)^{-1}(z'z)(x^{2'}z)^{-1}$ (Johnston,1984),  $s^2 = \frac{(y^*-b_2x^2)'(y^*-b_2x^2)}{n-2}$ ,  $Y^* = (Y - b_1X^0)$ ,  $y^* = Y^* - \bar{Y}^*$ 

Now we may test  $b_2$  for significance using student-t statistic. If found insignificant, we may restrict in (4), yielding

$$b_0 = \bar{Y} - b_1 \bar{X}^0 \tag{9}$$

and

$$\hat{X}_i = \frac{Y_i - b_0}{b_1}, \ i = 1, 2, \dots, n$$
 (10)

# 3 Recumbent Length

Until a baby can stand on his/her own feet by himself/herself properly, height of the baby is measured at lying position and this height is known as recumbent length in the literature of ergonomics. It is a very difficult job to measure this recumbent length at birth for lots of reasons. In general, this measurement is subject to errors. Growth monitoring of a baby upto one year of age is dependent on recumbent length at birth. Many other biostatistical and anthropological investigations depend on recumbent length at birth and thus demand a possible minimum error measure of recumbent length which we may refer as adjusted recumbent length. We can estimate the recumbent length at birth from various sources but our problem is to choose one that will be closer to the reality.

#### 3.1 Genesis

Rouche et. al. (1989) showed that the model

$$Y_t = a + b\sqrt{t} + c\,\log_e t + \epsilon_t \tag{11}$$

where  $Y_t$  is the recumbent length at time t and a, b, c are constants, fits the growth in recumbent length well in the first year of life if the measurements are taken in one month interval and can be used for prediction purpose. The major criticism of this model is that at time t = 0, recumbent length  $y_0$ , that is, recumbent length at birth becomes constant for all babies. Secondly, the model depends on the length of t. Pervin (l.c) collected primary data on 296 babies born between May 25-December 13, 2002 in different clinics and hospitals of Rajshahi City, Bangladesh and monitored their growth for 0-48 months. She (2004) fitted model (11) for recumbent length of 293 babies measured at 15 days interval at 23 points within the first year of life excluding the recumbent length at birth and found good fit for each and every baby separately. She then extended the growth curves backward to zero point freehand which cut the y-axis in some point above the origin. She considered that point as the recumbent length at birth. Since, the measures are based on graph, can not be used for farther statistical treatment. Moreover, ordinary least square fitted errors shows significant correlation with the estimated recumbent length indicating the presence of measurement errors in the estimated values.

#### 3.2 Model and Estimation

Height and weight of human being are well correlated. So is the recumbent length and weight at birth. Let  $Y_j$  be the weight in kg. and  $X_j$  be the recumbent length in cm. of the jth baby at birth. We may write the possible functional relationship as

(i) 
$$Y_j = \beta_0 + \beta_1 X_j + \epsilon_j$$
  
or  
(ii)  $Y_j = \beta_0 + \beta_1 X_j + \beta_2 X_j^2 + \epsilon_j$   
 $X_j^0 = X_j + u_j$   
 $j = 1, 2, \dots, n$ 
(12)

 $X_j$ 's being subject to measurement error, OLS can not be applied to the estimation of (12). To apply (4) and (12), we have to search for the instrumental variable z which is not related to Y or  $\epsilon$  but related to X only. We may define Z as below:

$$Z_j = Rank(X_j^0), \ j = 1, 2, \dots, n$$
 (13)

Pre-multiplying (i) of (12) by  $\frac{1}{n}Z'$  and applying assumption of (4), we have for  $Z = (z_1, z_2, \ldots, z_n)'$ ;  $Y = (y_1, y_2, \ldots, y_n)'$  and  $X^0 = (x_1^0, x_2^0, \ldots, x_n^0)'$ 

$$b_1 = \hat{\beta}_1 = \frac{\sum z_j y_j}{\sum z_j x_j^0}, \quad y_j = Y_j - \bar{Y}, \quad x_j^0 = X_j^0 - \bar{X}^0, \quad z_j = Z_j - \bar{Z}$$
(14)

 $\bar{X}^0$ ,  $\bar{Y}$ ,  $\bar{Z}$  being the arithmetic mean of  $X^0, Y, Z$ . Now we may construct  $Y_j^* = (Y_j - b_1 X_j^0)$  and regress  $Y_j^*$  on  $X_j^{0^2}$  using the same instrumental variable Z which was applied to (i) of (12) yielding the estimator

$$b_2 = \hat{\beta}_2 = \frac{\sum z_j y_j^*}{\sum z_j x_j^{0^2}} \sim AN\left(\beta_2, \frac{\sigma_u^2}{\left(\sum z_i x_j^{0^2}\right)^2}\right)$$
(15)

We can test  $b_2$  for significance in the OLS set up using Student's t-test. Finally we can obtain  $b_0 = \beta_0$  of (12) from (6) if  $b_2$  is found significant. Otherwise (9) may be used to obtain the same. Finally, we may have inverse estimates of  $X_j$  by using (8) or (10) as is the case may be.

## 4 **Results and Discussions**

Fitted model-1 of table-1 shows that both  $X^0$  and  $X^{0^2}$  are statistically significant. Also correlation between  $X^0$  and fitted errors are statistically significant. This implies that ordinary least square estimation of model-1 produces inconsistent estimates of parameters. Fitted model-2 of table-1 shows that instrumental variable estimators are consistent because correlation between fitted errors and the instrumental regressor is statistically insignificant. It also shows that quardetic term is insignificant so that the inverse regression estimator of X is provided by the equation (10). Fitted model-3 of table-1 with Pervin's (2004) estimate  $\hat{X}_1$  of X shows that correlation between  $\hat{X}_1$ and the corresponding fitted errors are statistically significant which restricts the use of  $\hat{X}_1$  as an estimate of X. Fitted model-4 of table-1 shows that inverse regression estimate  $X_2$  of X and the OLS fitted errors have statistically insignificant correlation. This implies that  $\hat{X}_2$  is free from measurement errors. But variance of  $\hat{X}_2$  is much larger than that of  $\hat{X}_1$ . Sum of squared distances between  $\hat{X}_2$  and  $X^0$  is also comparatively large than that of  $\hat{X}_1$  and  $\hat{X}^0$ . To reduce the variance and sum of squared distances, estimators  $\hat{X}_3 = \frac{\hat{X}_1 + \hat{X}_2}{2}$  and  $\hat{X}_4 = (\hat{X}_1 \cdot \hat{X}_2)^{\frac{1}{2}}$  are also proposed. Values of  $Y, X^0, \hat{X}_1, \hat{X}_2, \hat{X}_3$  and  $\hat{X}_4$  are appended in table-3. Fitted model-5 and 6 of table-1 shows that correlations between fitted errors and  $\hat{X}_3$  or  $\hat{X}_4$  are statistically insignificant which implies that both  $\hat{X}_3$  and  $\hat{X}_4$  are also measurement errors free estimates of X. Variances and sum of squared distances of  $\hat{X}_3$  and  $\hat{X}_4$  are also comparatively smaller than that of  $\hat{X}_2$ . Descriptive statistics of  $Y, X^0, \hat{X}_1, \hat{X}_2, \hat{X}_3$  and  $\hat{X}_4$  along with the sum of squared distance of  $\hat{X}_1, \hat{X}_2, \hat{X}_3$  and  $\hat{X}_4$  from  $X^0$  are appended in table-2.

Considering the ease of calculation, tolerable variation and sum of squared errors along with insignificant correlation with the fitted residuals, we recommend  $\hat{X}_3$  as the best estimator of X, the unknown recumbent length at birth.

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| Models  | Variables                                     | Coefficients | S. Error | t-value | Significance<br>Level | Correlation with fitted errors |
|---------|---|--------------|----------|---------|-----------------------|--------------------------------|
|         |   | OLS          |          |         | 20101                 |                                |
| Model 1 | Constant                                      | 23 403       | 8 1748   | 2 863   | 0.005                 | 0.000                          |
|         | $X^0$   | -0.926       | 0.3365   | -2.754  | 0.006                 | -0.858**                       |
|         | $X^{0^2}$                                     | 0.0104       | 0.0035   | 3.010   | 0.003                 | -                              |
|         |   | IV           |          |         |                       |                                |
| Model 2 | Constant                                      | -0.496       | 0.2357   | -2.101  | 0.003                 | 0.000                          |
|         | $Z'X^0$                                       | 0.0719       | 0.0236   | 3.450   | 0.006                 | 0.002                          |
|         | $Z'X^{0^2}$                                   | 0.0078       | 1.2195   | 0.006   | 0.958                 | -                              |
|         |   |              |          |         |                       |                                |
|         |   | OLS          |          |         |                       |                                |
| Model 3 | Constant                                      | 4.432        | 0.3031   | 14.622  | 0.000                 | 0.000                          |
|         | $\hat{X}_1$ (Pervin)                          | -0.069       | 0.0119   | -5.793  | 0.000                 | -0.840**                       |
|         | $\hat{X}_{1}^{2}$                             | 0.001        | 0.0001   | 6.824   | 0.000                 | -                              |
|         |   | OLS          |          |         |                       |                                |
| Model 4 | Constant                                      | -0.767       | 0.0256   | -30.171 | 0.000                 | 0.000                          |
|         | $\hat{X}_2$ (IRE)                             | 0.077        | 0.0008   | 96.250  | 0.000                 | 0.004                          |
|         |   | OLS          |          |         |                       |                                |
| Model 5 | Constant                                      | -2.962       | 0.1140   | -26.049 | 0.000                 | 0.000                          |
|         | $\hat{X}_3 = \frac{\hat{X}_{1+}\hat{X}_2}{2}$ | 0.122        | 0.2200   | 52.583  | 0.000                 | 0.017                          |
|         | £/  | OLS          |          |         |                       |                                |
| Model 6 | Constant                                      | -2.988       | 0.1140   | -26.190 | 0.000                 | 0.000                          |
|         | $\hat{X}_4 = (\hat{X}_1 + \hat{X}_2)^{1/2}$   | 0.123        | 0.0020   | 52.632  | 0.000                 | 0.015                          |
|         | l ` ´   |              |          |         |                       |                                |

Table 1: Fitted models and Test Results

OLS: Ordinary Least Squares IV: Instrumental Variable

IRE: Inverse Regression Estimator

\*\* Correlation is significant at the 0.01 level (2-tailed).

Sum of Squared distance from  $X^0$ Variables Mean Variance Ν Υ 296 2.999 0.241\_  $\mathbf{X}^{0}$ 29648.6255.889- $\hat{X}_1$ 29648.5456.35345.7711673.98 $\hat{X}_2$ 29648.62546.605 $\hat{X}_3$ 16.25329648.5852842.66 $\hat{X}_4$ 29648.48216.0752832.40

Table 2: Descriptive Statistics and Sum of Squared distance from  $X^0$ 

| Table 5. Observed and Estimated values of weights and flecombent Dength at Dirth | Table 3: | Observed | and | Estimated | Values o | of Weig | hts and | Recumbent | Length | $\mathbf{at}$ | Birth |
|--|----------|----------|-----|-----------|----------|---------|---------|-----------|--------|---------------|-------|
|--|----------|----------|-----|-----------|----------|---------|---------|-----------|--------|---------------|-------|

| ST  | V  | <b>v</b> <sup>0</sup>  | Ŷ.   | Ŷ-   | Ŷ-  | Ŷ.   |
|---|--|--|--|--|---|--|
| 51  | 1  | A  | <u> </u>   | A2   | A3  | A4   |
|   | 2.30   | 42.00  | 42.00  | 38.90  | 40.45   | 40.42  |
| 2   | 2.30   | 43.00  | 42.00  | 38.90  | 40.45   | 40.42  |
| 3   | 2.70   | 43.00  | 43.00  | 44.46  | 43.73   | 43.73  |
| 4   | 3 50   | 44.00  | 43.00  | 55.60  | 49.30   | 48.89  |
|   | 0.00   | 44.00  | 40.00  | 27 51  | 40.75   | 40.60  |
| 5   | 2.20   | 44.00  | 44.00  | 37.51  | 40.75   | 40.02  |
| 6   | 2.50   | 44.00  | 44.00  | 41.68  | 42.84   | 42.82  |
| 7   | 2.70   | 44.00  | 44.00  | 44.46  | 44.23   | 44.23  |
| 8   | 2.50   | 44.00  | 44.00  | 41.68  | 42.84   | 42.82  |
| 0   | 2.50   | 44.00  | 44.00  | 41.69  | 42.84   | 42.92  |
| 10  | 2.00   | 44.00  | 44.00  | 41.00  | 42.04   | 42.02  |
| 10  | 3.00   | 44.00  | 44.00  | 48.64  | 46.32   | 46.26  |
| 11  | 2.30   | 44.00  | 44.00  | 38.90  | 41.45   | 41.37  |
| 12  | 2.70   | 44.00  | 44.40  | 44.46  | 44.43   | 44.43  |
| 13  | 2.60   | 45.00  | 44.00  | 43.07  | 43.54   | 43.53  |
| 14  | 3 50   | 45.00  | 44.00  | 55.60  | 49.80   | 49.46  |
| 15  | 0.00   | 45.00  | 44.00  | 41.69  | 49.00   | 40.40  |
| 15  | 2.50   | 45.00  | 44.00  | 41.08  | 42.84   | 42.82  |
| 16  | 2.70   | 45.00  | 44.40  | 44.46  | 44.43   | 44.43  |
| 17  | 2.00   | 45.00  | 45.00  | 34.72  | 39.86   | 39.53  |
| 18  | 2.70   | 45.00  | 45.00  | 44.46  | 44.73   | 44.73  |
| 19  | 3 50   | 45.00  | 44.00  | 55.60  | 49.80   | 49.46  |
| 20  | 2 70   | 45.00  | 45.00  | 14 46  | 44 79   | 44 79  |
| 20  | 4.70   | 40.00  | 45.00  | 44.40  | 44.73   | 44.73  |
| 21  | 2.50   | 45.00  | 45.00  | 41.68  | 43.34   | 43.31  |
| 22  | 2.50   | 45.00  | 45.00  | 41.68  | 43.34   | 43.31  |
| 23  | 3.50   | 45.00  | 45.00  | 55.60  | 50.30   | 50.02  |
| 24  | 2.70   | 45.00  | 45.00  | 44.46  | 44.73   | 44.73  |
| 25  | 3.00   | 45.00  | 45.00  | 48.64  | 46.82   | 46.78  |
| 20  | 0.00   | 45.00  | 45.00  | 40.04  | 44.72   | 44.72  |
| 20  | 2.70   | 40.00  | 40.00  | 44.40  | 44./3   | 44.73  |
| 27  | 2.50   | 45.00  | 45.00  | 41.68  | 43.34   | 43.31  |
| 28  | 3.00   | 45.00  | 45.00  | 48.64  | 46.82   | 46.78  |
| 29  | 2.60   | 45.00  | 45.00  | 43.07  | 44.04   | 44.03  |
| 30  | 2.50   | 45.00  | 45.00  | 41.68  | 43.34   | 43.31  |
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| $ \begin{array}{c} \dots \\ \dots \\ \dots \\ \dots \\ 267 \\ 268 \\ 269 \\ \end{array} $   | <br><br>2.50<br>2.50<br>2.50   | ····<br>···<br>52.00<br>52.00<br>52.00   | ····<br>····<br>52.00<br>52.00<br>52.00  | ····<br>···<br>41.68<br>41.68<br>41.68   | ····<br>···<br>46.84<br>46.84   |  |
| $ \begin{array}{c} \dots \\ \dots \\ \dots \\ 267 \\ 268 \\ 269 \\ 270 \\ \end{array} $   | ····<br>···<br>2.50<br>2.50<br>2.50<br>2.50  | ····<br>····<br>52.00<br>52.00<br>52.00<br>52.00   | <br><br>52.00<br>52.00<br>52.00<br>52.00   | <br><br>41.68<br>41.68<br>41.68<br>41.68   | <br><br>46.84<br>46.84<br>46.84<br>46.84  | <br><br>46.55<br>46.55<br>46.55<br>46.55   |
| $ \begin{array}{c} \dots \\ \dots \\ \dots \\ 267 \\ 268 \\ 269 \\ 270 \\ 271 \\ \end{array} $  | ····<br>···<br>2.50<br>2.50<br>2.50<br>2.50<br>2.50<br>3.57  | <br>52.00<br>52.00<br>52.00<br>52.00   | ····<br>···<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00   | $\begin{array}{c} \dots \\ \dots \\ \dots \\ \dots \\ \dots \\ \dots \\ 1.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 56.57 \end{array}$   | <br><br>46.84<br>46.84<br>46.84<br>46.84<br>46.84   | <br><br>46.55<br>46.55<br>46.55<br>46.55<br>53.71  |
| $ \begin{array}{c} \dots \\ \dots \\ \dots \\ 267 \\ 268 \\ 269 \\ 270 \\ 271 \\ 272 \\ \end{array} $   | <br>2.50<br>2.50<br>2.50<br>2.50<br>3.57<br>2.50   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00   | 52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00   | $\begin{array}{c} & & \\$ | <br><br>46.84<br>46.84<br>46.84<br>46.84<br>46.84<br>53.78  | <br><br>46.55<br>46.55<br>46.55<br>46.55<br>53.71  |
| $ \begin{array}{c} \dots \\ \dots \\ 267 \\ 268 \\ 269 \\ 270 \\ 271 \\ 272 \\ 27$      | <br>2.50<br>2.50<br>2.50<br>2.50<br>3.57<br>2.50   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00   | $\begin{array}{c} & & \\$ | <br><br>46.84<br>46.84<br>46.84<br>46.84<br>53.78<br>46.84  | $\begin{array}{c} & & \\$ |
| $\begin{array}{c} \dots \\ \dots \\ \dots \\ 267 \\ 268 \\ 269 \\ 270 \\ 271 \\ 272 \\ 273 \\ 273 \end{array}$  | 2.50           2.50           2.50           2.50           3.57           2.50           3.20   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00   | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 1.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 56.57 \\ 41.68 \\ 51.42 \\ \end{array}$   | $\begin{array}{c} \dots \\ & \\ & \\ & \\ & \\ & \\ & \\ & \\ & \\ & \\$  | $\begin{array}{c} \dots \\ & \\ & \\ & \\ & \\ & \\ & \\ & \\ & \\ & \\$   |
| $\begin{array}{c} \dots \\ \dots \\ \dots \\ 267 \\ 268 \\ 269 \\ 270 \\ 271 \\ 272 \\ 273 \\ 274 \end{array}$  | <br>2.50<br>2.50<br>2.50<br>3.57<br>2.50<br>3.20<br>3.20   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | $\begin{array}{c} \dots \\ \dots \\ 41.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 56.57 \\ 41.68 \\ 51.42 \\ 48.64 \end{array}$  | $\begin{array}{c} \dots \\ & \\ & \\ & \\ & \\ & \\ & \\ & \\ & \\ & \\$  | $\begin{array}{c} \dots \\ & \\ & \\ & \\ & \\ & \\ & \\ & \\ & \\ & \\$   |
| $\begin{array}{c} \dots \\ \dots \\ 267 \\ 268 \\ 269 \\ 270 \\ 271 \\ 272 \\ 273 \\ 274 \\ 275 \end{array}$  | <br>2.50<br>2.50<br>2.50<br>3.57<br>2.50<br>3.20<br>3.20<br>3.00<br>2.70   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00   | $\begin{array}{c} \dots \\ \dots $   | $\begin{array}{c} \dots \\ \dots \\ 46.84 \\ 46.84 \\ 46.84 \\ 46.84 \\ 53.78 \\ 46.84 \\ 51.71 \\ 50.32 \\ 48.23 \end{array}$  | $\begin{array}{c} \dots \\ & \\ & \\ & \\ & \\ & \\ & \\ & \\ & \\ & \\$   |
| $\begin{array}{c} \dots \\ \dots \\ \dots \\ 267 \\ 268 \\ 269 \\ 270 \\ 271 \\ 272 \\ 273 \\ 274 \\ 275 \\ 276 \end{array}$  | <br>2.50<br>2.50<br>2.50<br>3.57<br>2.50<br>3.20<br>3.20<br>3.20<br>3.40   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | $\begin{array}{c} \dots \\ & \dots \\ & \dots \\ & & 1.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 56.57 \\ 41.68 \\ 51.42 \\ 48.64 \\ 44.46 \\ 54.20 \end{array}$  | $\begin{array}{c} \dots \\ 46.84 \\ 46.84 \\ 46.84 \\ 46.84 \\ 53.78 \\ 46.84 \\ 51.71 \\ 50.32 \\ 48.23 \\ 53.10 \end{array}$  | $\begin{array}{c} & & & \\$  |
| $\begin{array}{c} \cdots \\ \hline \cdots \\ \hline 267 \\ \hline 268 \\ \hline 269 \\ \hline 270 \\ \hline 271 \\ \hline 272 \\ \hline 273 \\ \hline 274 \\ \hline 275 \\ \hline 276 \\ \hline 277 \\ \hline \end{array}$  | <br>2.50<br>2.50<br>2.50<br>2.50<br>3.57<br>3.20<br>3.00<br>2.70<br>3.00<br>2.70<br>3.00   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00   | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 1.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 56.57 \\ 41.68 \\ 51.42 \\ 48.64 \\ 44.46 \\ 54.20 \\ 34.72 \end{array}$   | $\begin{array}{c} \dots \\ \dots \\ \dots \\ \dots \\ 46.84 \\ 46.84 \\ 46.84 \\ 46.84 \\ 53.78 \\ 46.84 \\ 51.71 \\ 50.32 \\ 48.23 \\ 53.10 \\ 43.36 \end{array}$  | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 46.55 \\ 46.55 \\ 46.55 \\ 53.71 \\ 46.55 \\ 51.71 \\ 50.29 \\ 48.08 \\ 53.09 \\ 42.49 \end{array}$   |
| $\begin{array}{c} \dots \\ \dots \\ \dots \\ 267 \\ 268 \\ 269 \\ 270 \\ 271 \\ 272 \\ 273 \\ 274 \\ 275 \\ 276 \\ 277 \\ 278 \end{array}$  | <br>2.50<br>2.50<br>2.50<br>2.50<br>3.57<br>2.50<br>3.20<br>3.00<br>2.70<br>3.40<br>2.00<br>3.50   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | $\begin{array}{c} \dots \\ \dots \\ 1 \\ 1 \\ 1 \\ 1 \\ 1 \\ 1 \\ 1 \\ 1 \\$   | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 46.84 \\ 46.84 \\ 46.84 \\ 46.84 \\ 53.78 \\ 46.84 \\ 53.78 \\ 46.84 \\ 53.71 \\ 50.32 \\ 48.23 \\ 53.10 \\ 43.36 \\ 53.80 \end{array}$  | <br>46.55<br>46.55<br>46.55<br>53.71<br>46.55<br>53.71<br>46.55<br>51.71<br>50.29<br>48.08<br>53.09<br>48.08<br>53.09  |
| $\begin{array}{c} \dots \\ \dots \\ 267 \\ 268 \\ 269 \\ 270 \\ 271 \\ 272 \\ 273 \\ 274 \\ 275 \\ 276 \\ 277$      | <br>2.50<br>2.50<br>2.50<br>2.50<br>3.57<br>2.50<br>3.20<br>3.20<br>3.00<br>2.70<br>3.40<br>2.00<br>3.40<br>2.00   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00   | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 1.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 56.57 \\ 41.68 \\ 51.42 \\ 48.64 \\ 44.66 \\ 54.20 \\ 34.72 \\ 55.60 \\ 14.20 \\ 55.60 \\ 14.2$   | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 46.84 \\ 46.84 \\ 46.84 \\ 46.84 \\ 46.84 \\ 46.84 \\ 53.78 \\ 46.84 \\ 53.70 \\ 31.0 \\ 53.30 \\ 53.30 \\ 53.8$      | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 46.55 \\ 46.55 \\ 46.55 \\ 46.55 \\ 53.71 \\ 46.55 \\ 51.71 \\ 50.29 \\ 48.08 \\ 53.09 \\ 42.49 \\ 53.77 \\ 52.$   |
| $\begin{array}{c} \cdots \\ \cdots \\ 267 \\ 268 \\ 269 \\ 270 \\ 271 \\ 272 \\ 273 \\ 274 \\ 275 \\ 276 \\ 277 \\ 278 \\ 277 \\ 278 \\ 279 \\ 26 \\ 26 \\ 26 \\ 26 \\ 27 \\ 278 \\ 279 \\ 279 \\ 278 \\ 279 \\ 270 \\ 27$   | <br>2.50<br>2.50<br>2.50<br>2.50<br>3.57<br>2.50<br>3.20<br>3.00<br>2.70<br>3.00<br>2.70<br>3.00<br>2.00<br>3.50<br>3.50<br>3.50   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00   | $\begin{array}{c} \dots \\ \dots \\ 1.11111111111111111111111111111$   | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 46.84 \\ 46.84 \\ 46.84 \\ 46.84 \\ 53.78 \\ 46.84 \\ 53.78 \\ 51.71 \\ 50.32 \\ 48.23 \\ 51.70 \\ 43.36 \\ 53.80 \\ 53.$      | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 46.55 \\ 46.55 \\ 46.55 \\ 53.71 \\ 46.55 \\ 51.71 \\ 46.55 \\ 51.71 \\ 51.71 \\ 51.29 \\ 48.08 \\ 53.09 \\ 42.49 \\ 53.77 \\ 53.77 \\ 53.77 \\ 53.77 \\ 53.77 \end{array}$   |
| $\begin{array}{c} \hline \cdots \\ \hline \cdots \\ \hline 267 \\ 268 \\ 269 \\ 270 \\ 271 \\ 272 \\ 273 \\ 274 \\ 275 \\ 276 \\ 277 \\ 278 \\ 279 \\ 280 \\ \end{array}$   | <br>2.50<br>2.50<br>2.50<br>3.57<br>2.50<br>3.20<br>3.20<br>3.00<br>2.70<br>3.40<br>2.70<br>3.40<br>2.50<br>3.50<br>3.50<br>3.20   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | $\begin{array}{c} \dots \\ \dots \\ 1.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 56.57 \\ 41.68 \\ 51.42 \\ 48.64 \\ 44.46 \\ 54.20 \\ 34.72 \\ 55.60 \\ 51.42 \end{array}$  | $\begin{array}{c} \dots \\ \dots $  | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 46.55 \\ 46.55 \\ 46.55 \\ 53.71 \\ 46.55 \\ 51.71 \\ 50.29 \\ 48.08 \\ 53.09 \\ 42.49 \\ 53.77 \\ 53.77 \\ 51.71 \end{array}$  |
| $\begin{array}{c} \ddots \\ \cdots \\ 267\\ 268\\ 269\\ 270\\ 271\\ 272\\ 273\\ 274\\ 275\\ 276\\ 277\\ 278\\ 277\\ 278\\ 279\\ 280\\ 281\\ \end{array}$  | <br>2.50<br>2.50<br>2.50<br>2.50<br>3.57<br>2.50<br>3.20<br>3.00<br>2.70<br>3.20<br>3.20<br>3.20<br>3.20<br>3.50<br>3.50<br>3.50<br>3.20   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00   | $\begin{array}{c} \dots \\ \dots \\ 11.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 56.57 \\ 41.68 \\ 51.42 \\ 48.64 \\ 44.46 \\ 51.42 \\ 55.60 \\ 54.20 \\ 55.60 \\ 55.60 \\ 55.60 \\ 55.60 \\ \end{array}$   | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 46.84 \\ 46.84 \\ 46.84 \\ 46.84 \\ 46.84 \\ 46.84 \\ 53.78 \\ 46.84 \\ 53.78 \\ 53.78 \\ 53.78 \\ 53.10 \\ 53.80 \\ 53.80 \\ 53.80 \\ 53.80 \\ 51.71 \\ 53.80 \end{array}$  | $\begin{array}{c} \dots \\ \dots \\ 46.55 \\ 46.55 \\ 46.55 \\ 46.55 \\ 46.55 \\ 53.71 \\ 46.55 \\ 51.71 \\ 50.29 \\ 48.08 \\ 53.09 \\ 53.77 \\ 53.77 \\ 53.77 \\ 53.77 \\ 53.77 \end{array}$  |
| $\begin{array}{c} \ddots \\ \cdots \\ 267 \\ 268 \\ 269 \\ 270 \\ 271 \\ 272 \\ 273 \\ 274 \\ 275 \\ 276 \\ 277 \\ 276 \\ 277 \\ 278 \\ 279 \\ 280 \\ 281 \\ 282 \end{array}$   | $\begin{array}{c} \dots \\ \dots \\ 2.50 \\ 2.50 \\ 2.50 \\ 3.57 \\ 2.50 \\ 3.20 \\ 3.20 \\ 3.20 \\ 3.40 \\ 2.70 \\ 3.40 \\ 3.50 \\ 3.50 \\ 3.50 \\ 3.50 \\ 3.20 \\ 3.40 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\$  | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | $\begin{array}{c} \dots \\ \dots \\ 1.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 56.57 \\ 41.68 \\ 51.42 \\ 48.64 \\ 54.20 \\ 34.72 \\ 55.60 \\ 55.60 \\ 55.60 \\ 51.42 \\ 51.42 \\ $   | $\begin{array}{c} \dots \\ \dots $  | $\begin{array}{c} \dots \\ & 46.55 \\ & 46.55 \\ & 46.55 \\ & 53.71 \\ & 50.29 \\ & 46.55 \\ & 53.77 \\ & 53.77 \\ & 53.77 \\ & 53.77 \\ & 53.77 \\ & 53.77 \\ & 53.77 \\ & 51.71 \\ & 53.77 \\ & 57.03 \end{array}$   |
| $\begin{array}{c} \dots \\ \dots \\ 267 \\ 268 \\ 269 \\ 270 \\ 271 \\ 272 \\ 273 \\ 274 \\ 275 \\ 276 \\ 277 \\ 278 \\ 279 \\ 280 \\ 281 \\ 282 \\ 283 \end{array}$  | <br>2.50<br>2.50<br>2.50<br>3.57<br>2.50<br>3.20<br>3.00<br>2.70<br>3.40<br>2.70<br>3.40<br>2.00<br>3.50<br>3.50<br>3.50<br>4.00   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | $\begin{array}{c} \dots \\ \dots \\ 1.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 56.57 \\ 41.68 \\ 51.42 \\ 48.64 \\ 44.46 \\ 51.42 \\ 34.72 \\ 55.60 \\ 55.60 \\ 51.42 \\ 55.60 \\ 51.42 \\ 55.60 \\ 62.55 \\ 69.51 \end{array}$  | $\begin{array}{c} \dots \\ \dots \\ \dots \\ \dots \\ 46.84 \\ 46.84 \\ 46.84 \\ 46.84 \\ 53.70 \\ 46.84 \\ 51.71 \\ 50.32 \\ 48.23 \\ 53.10 \\ 43.36 \\ 53.80 \\ 51.71 \\ 53.80 \\$ | $\begin{array}{c} \dots \\ \dots \\ \dots \\ \dots \\ 46.55 \\ 46.55 \\ 46.55 \\ 46.55 \\ 53.71 \\ 46.55 \\ 51.71 \\ 50.29 \\ 48.08 \\ 53.09 \\ 42.49 \\ 53.77 \\ 53.77 \\ 51.71 \\ 53.77 \\ 51.71 \\ 53.77 \\ 51.70 \\ 60.12 \end{array}$   |
| $\begin{array}{c} \dots \\ \dots \\ \dots \\ 267 \\ 268 \\ 269 \\ 270 \\ 271 \\ 272 \\ 273 \\ 274 \\ 275 \\ 276 \\ 277 \\ 278 \\ 279 \\ 280 \\ 281 \\ 282 \\ 283 \\ 283 \\ 284 \\ \end{array}$  | <br>2.50<br>2.50<br>2.50<br>2.50<br>2.50<br>3.57<br>2.50<br>3.20<br>3.00<br>2.70<br>3.20<br>3.00<br>2.70<br>3.50<br>3.50<br>3.50<br>3.20<br>3.50<br>4.00<br>4.00   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 141.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 56.57 \\ 41.68 \\ 51.42 \\ 48.64 \\ 44.46 \\ 51.42 \\ 55.60 \\ 55.60 \\ 55.60 \\ 51.42 \\ 55.60 \\ 62.55 \\ 69.51 \\ 42.55 \\ \end{array}$  | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 46.84 \\ 46.84 \\ 46.84 \\ 46.84 \\ 46.84 \\ 453.78 \\ 46.84 \\ 53.78 \\ 53.78 \\ 53.78 \\ 53.80 \\ 53.80 \\ 53.80 \\ 51.71 \\ 53.80 \\ 51.72 \\ 83.80 \\ 57.28 \\ 60.76 \\ 57.28 \\ \end{array}$  | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 46.55 \\ 46.55 \\ 46.55 \\ 46.55 \\ 46.55 \\ 53.71 \\ 46.55 \\ 51.71 \\ 50.29 \\ 48.08 \\ 53.09 \\ 42.49 \\ 53.77 \\ 53.77 \\ 53.77 \\ 51.71 \\ 53.77 \\ 51.71 \\ 53.77 \\ 57.03 \\ 60.12 \\ 57.03 \end{array}$   |
| $\begin{array}{c} \dots \\ \dots \\ 267\\ 268\\ 269\\ 270\\ 271\\ 272\\ 273\\ 274\\ 275\\ 276\\ 277\\ 278\\ 279\\ 280\\ 281\\ 282\\ 283\\ 284\\ 283\\ 284\\ 283\\ 284\\ 285\\ 285\\ 285\\ 285\\ 285\\ 285\\ 285\\ 285$  | <br>2.50<br>2.50<br>2.50<br>3.50<br>3.20<br>3.20<br>3.20<br>3.20<br>3.20<br>3.40<br>2.70<br>3.40<br>2.70<br>3.40<br>2.70<br>3.50<br>3.50<br>3.50<br>3.50<br>3.50<br>4.00<br>4.50   | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 1.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 56.57 \\ 41.68 \\ 51.42 \\ 48.64 \\ 44.46 \\ 54.20 \\ 34.72 \\ 55.60 \\ 51.42 \\ 55.60 \\ 51.42 \\ 55.60 \\ 51.42 \\ 55.60 \\ 51.62 \\ 55.69 \\ 51.62 \\ 55.60 \\ 52.55 \\ 69.51 \\ 62.55 \\ 62.5$   | $\begin{array}{c} \dots \\ \dots $  | $\begin{array}{c} \dots \\ \dots $   |
| $\begin{array}{c} \dots \\ \dots \\ 267\\ 268\\ 269\\ 270\\ 271\\ 272\\ 273\\ 274\\ 275\\ 276\\ 277\\ 278\\ 276\\ 277\\ 278\\ 279\\ 280\\ 281\\ 282\\ 283\\ 284\\ 282\\ 283\\ 284\\ 285\\ 284\\ 285\\ 284\\ 285\\ 284\\ 285\\ 284\\ 285\\ 284\\ 285\\ 284\\ 285\\ 284\\ 285\\ 286\\ 286\\ 286\\ 286\\ 286\\ 286\\ 286\\ 286$  | $\begin{array}{c} \dots \\ \dots \\ 2.50 \\ 2.50 \\ 2.50 \\ 2.50 \\ 2.50 \\ 3.20 \\ 3.20 \\ 3.00 \\ 2.70 \\ 3.20 \\ 3.20 \\ 3.20 \\ 3.20 \\ 3.50 \\ 3.50 \\ 3.50 \\ 3.50 \\ 3.50 \\ 4.00 \\ 4.00 \\ 4.00 \\ 4.00 \\ 4.00 \\ 2.0 \\ \end{array}$  | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | <br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00<br>52.00  | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 11.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 41.68 \\ 56.57 \\ 41.68 \\ 51.42 \\ 48.64 \\ 44.46 \\ 51.42 \\ 55.60 \\ 34.72 \\ 55.60 \\ 55.60 \\ 51.42 \\ 55.60 \\ 51.42 \\ 55.60 \\ 62.55 \\ 69.51 \\ 62.55 \\ 62.$   | <br>46.84<br>46.84<br>46.84<br>46.84<br>46.84<br>46.84<br>46.84<br>46.84<br>46.84<br>46.84<br>46.84<br>46.84<br>46.84<br>46.84<br>46.84<br>453.71<br>50.32<br>48.23<br>53.10<br>43.36<br>53.80<br>53.80<br>51.71<br>53.80<br>51.71<br>53.80<br>57.28<br>60.76<br>57.28  | $\begin{array}{c} \dots \\ \dots \\ \dots \\ 46.55 \\ 46.55 \\ 46.55 \\ 46.55 \\ 46.55 \\ 53.71 \\ 46.55 \\ 51.71 \\ 50.29 \\ 48.08 \\ 53.09 \\ 42.49 \\ 53.77 \\ 53.77 \\ 53.77 \\ 51.71 \\ 53.77 \\ 51.71 \\ 53.77 \\ 51.703 \\ 60.12 \\ 57.03 \\ 57.03 \\ 52.20 \end{array}$  |
| $\begin{array}{c} \dots \\ \dots \\ 267\\ 268\\ 269\\ 270\\ 271\\ 272\\ 273\\ 274\\ 275\\ 276\\ 277\\ 278\\ 279\\ 280\\ 281\\ 282\\ 283\\ 284\\ 285\\ 285\\ 285\\ 285\\ 285\\ 285\\ 285\\ 285$  | <br>2.50<br>2.50<br>2.50<br>3.57<br>3.20<br>3.20<br>3.20<br>3.20<br>3.40<br>2.70<br>3.40<br>2.70<br>3.40<br>2.70<br>3.50<br>3.50<br>3.50<br>3.50<br>3.50<br>3.50<br>3.50<br>3.5  | 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\\ 60.65 \\ 60.65 \\ 59.95 \\ 59.30 \end{array}$   |